THE ISOLATION OF SUB-SAHARAN FLOATING CURRENCIES

[Izolace subsaharských plovoucích měnových kurzů]

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Abstract: The paper tested contagion effects among free floating African currencies and Euro with a control variable like the Euro-pegged CFA Franc. Contagion should be based on deep trade, funding and political relationships which was valid for connections between African countries and such developed countries like the United States or the Eurozone. The theory suggests increased common movements under shock periods which were tested on a daily time series between 2000 and 2015, studying relationships under recession periods in the United States or the Eurozone or under days with extreme fluctuation. The results presented contagions only for the emerging South African currency while the others proved to be relatively independent.

Keywords: Africa, contagion, extreme fluctuation, floating regime.

JEL classification: C32, F31, F35

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Introduction

Current paper evaluates the consequences of floating on Sub-Saharan African currencies to see the possible trade-offs between price-competitiveness and foreign exchange (FX) exposure. Peg is a political commitment to eliminate FX exposure with a perfect comovement between pegged and amplitude of extreme fluctuation is limited. However, pricecompetitiveness can be biased by appreciating trends of the selected key currency – while the maintenance can be expensive under turbulent times as well. Floating regime has the advantage of adaption in our current two-key-currency world, but later defined contagions and divergences can ruin the profitability due to unexpected change in currency common movements. The operative objective of this paper is to check upper hypotheses about currency regimes on the selected country set: occurrence of extreme currency fluctuations should be limited under peg, they should be isolated from volatility spillover and shall maintain a strong correlation with Euro (IMF 2013). Floating regimes should present the well-known properties increased extreme fluctuation under crisis periods, some sort of volatility spillover with neighbouring countries and other key trade partners as well as the strength of the comovement with Euro should be the subject of market conditions. These results can be important to calibrate international and national aid and development programs.

After the definition of contagion and its background are discussed, special African tradeprograms will be summarized in the theoretical chapter. Current study tested daily closing data of floating currencies from Kenya (KES), Ghana (GHS), South Africa (ZAR), Tanzania (TZS), Uganda (UGX), Gambia (GMD), Madagascar (MGA) and Mozambique (MZN) in USD denomination against EUR/USD rate between March 8 2000 and March 6 2015 acquired from Bloomberg database. These currencies have poor literature despite the enormous international efforts to stabilize these countries external imbalances. The selected time period covers three crisis periods (dot-com, subprime and PIGS), allowing us to capture currency misbehaviours under extreme circumstances. The basic statistics and the ways to analyse their extreme fluctuation, volatility and correlation are the content of the data and methods chapter. Results are supporting the advantages of floating regime, which was remarkably isolated from key currency developments.

1 Theoretical Background

This chapter defines contagion channels to study vulnerability of sample countries. Floating exchange rate regime was applied in the selected countries, while their external balance is in focus of the presented international initiatives in the following paragraphs.

Contagions could be broadly defined as the cross-country transmission of shocks or the general cross-country spillover effects, which does not need to be related to crises. Current paper applies the World Bank's very restrictive definition¹: a relative increase² in cross-country correlations during "crisis times" to "tranquil times" - when a shock to one or a group of markets, countries, or institutions, spread to other markets, or countries, or institutions (Pritsker, 2001). Contagion is based on three fundamental links among countries, like financial, real and political links. Financial link is supported by connections through the international financial system (common financial institutions, interconnected lenders, nonbank financial market participants). Real links are in connection with international trade or FDI-driven cross-border division of labour. Political link is in connection with exchange rate arrangement country-groups as well as other current transfers or debt forgiveness due to international cooperation. Contagion transmission channels can be paired with specific balance of payments components which will be used during further evaluations of Africandeveloped-emerging relationships.

The analyzed African country set followed free floating exchange rate mechanism combined with monetary aggregate target in the majority of the cases (Kenya, Tanzania, Uganda, Gambia, Madagascar, Mozambique) or inflation targeting framework (Ghana, South Africa) (IMF, 2013). Main trading partners were the European Union, United States, China and India in 2012 (CIA, 2015). Foreign trade was supported via a rich and chaotic net of regional integration agreements fostering trade of goods and services among selected countries and key economic areas (EU, US, Indian and Chinese markets) (Udvari 2012). Aid for Trade (AfT) initiative was created after the G8 Summit in 2005 to improve supply-side capacities of recipient countries instead of former "trade not aid" philosophy. The objective of this program was to integrate the recipient countries in the world economy, to diversify their foreign trade and to maintain their price-competitiveness under the current environment with decreasing tariffs - where former Generalized System of Preferences (GSP) became obsolete (Udvari, 2011). The European Union has been focusing most of its international aid-activity on African countries since the first Lome Agreement in 1975 even in the current Cotonou Agreement between 2000 and 2020 trough its financial (European Development Fund, dedicated EIB credits, economic stability aid programs) and trade instruments (GSP, unilateral preferences, WTO agreements and European Partnership Agreements) (Udvari, 2008, 2011, 2012). The U.S. also offers incentives for Sub-Saharan countries to open their economies and build free markets. The African Growth and Opportunity Act (AGOA) was approved by the U.S government in 2000. AGOA is a nonreciprocal trade preference program that provides dutyfree treatment to U.S. imports of certain products from actually 39 SSA countries; the Act is similar to the GSP but includes trade and development provisions beyond duty-free

¹ see: http://go.worldbank.org/JIBDRK3YC0

 $^{^{2}}$ We can talk about interdependence, when the upper difference between correlations under extreme and normal conditions is insignificant – meaning that upper links have no significant impact on exchange rates.

preferences (Williams, 2015). Project financing was a traditional field of World Bank Group, especially International Development Association (IDA) to provide interest-free loans for governments of Heavily Indebted Poor Countries (HIPC). Funding for corporate sector has a bigger variety: among another WB subsidiaries and regional development banks, the China-Africa Development Fund was established in 2007 to finance development projects on commercial basis for corporate sector, supported by China Development Bank (CADF, 2013). Thanks to South-South economic engagements, in recent years India and China as emerging economies have also become the main trade partners for African economies in a range of products (Dash, 2014). Upper programs are common in their political influence on target countries current and financial accounts, enriching the range of contagion possibilities (Fig. 1).

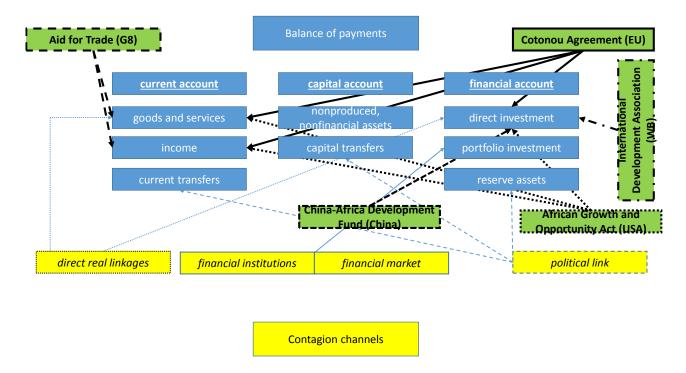


Figure 1: Contagion Channels and Balance of Payments

Source: IMF (2009), Pritsker (2001)

However, it can be misleading to treat Sub-Saharan Africa as a club due to the enormous intra and inter-regional deviations, causing lack of convergence among them (Mccoskey, 2002). Despite the upper presented wide range of support channels, Lomé and Cotonou Agreements had no significant impact on sigma (income deviations), beta (less developed has bigger growth-potential) and stochastic (entire country-group develops among a common trend) convergence (Gáspár and Udvari, 2011) or trade with European countries (Udvari, 2014). GDP per capita growth in Africa was increased in the past mostly by enhancements in literacy, investment, FDI and urban agglomeration via strengthening human and physical capital (Naudé, 2004). These results can make political and real linkages as primary contagion channels – because all of the political initiatives were made to compensate insufficient financial conditions. Can we increase the resilience of African countries with weak and policy-driven financial links? Past crisis experience supports this idea: South African stock exchange and bond markets suggested a marginal depreciation during the Asian crisis (Siriwardana and Dollery, 2002), but under global risk aversion periods a wide range of developed and emerging currencies (like South African Rand) tended to depreciate (De Bock and de Carvalho Filho, 2015). Size, liquidity and access are influencing factors of foreign investor participation to integrate a market into global capital flows – Collins and Biekpe (2003) suggest that as a reason why only stock markets in South Africa, Egypt, Namibia and Morocco presented contagion under Asian crisis. This idea is supported by Morales and Andreosso-O'Callaghan (2014) who suggest that financial crisis impact related to market integration level.

Current paper cites Visegrad group³-related experience to focus expectations. This country-set has been the subject of numerous articles to study their affections for contagions – while they are also the subject of policy-related transfers and their FDI, trade and banking relations are dominated by Eurozone-located factors but the overall size of their banking system or capital market is marginal (Farkas, 2011, Heryán, 2015). These relations had significant impact on strong common movement among currencies (Stavárek, 2010, Babetskaia-Kukharchuk et al., 2008, Bubák et al., 2011, Kiss and Schuszter, 2015), biased by crisis-related developments. Obviously these countries are closer to the developed than to the emerging status, while only South Africa is recognized as an emerging market from our African country-set according to their BRICS membership⁴. However, it can underline some significant differences among poor and emerging currencies – like the tightness of their common movements, volatility spillovers and extreme fluctuations.

Foreign exchange (FX) exposure can have a crucial impact on the competitiveness of analyzed African countries due to their floating currency regime framework. Combined with high dependence on euro and US dollar denominated markets (China is considered here, because of the 2% floating band of RMB against US dollar⁵) and diverse financial and trade support channels, contagion on currency market is a real option. This is the opposite strategy than that followed by member states⁶ of West African Economic and Monetary Union⁷, where CFA franc has a fixed parity to the Euro (656 to 1).

2 Data and Methods

Methods to capture temporal distribution of extreme FX fluctuations and contagions under recession periods in developed markets are presented in this chapter. The literature of African currencies is really narrow: only South African Rand was studied by De Bock et al. (2015) or Keskina et al. (2011) while Collins and Biekpe (2003) or Siriwardana and Dollery (2002) studied a wider African subset of equity markets with a result of isolation.

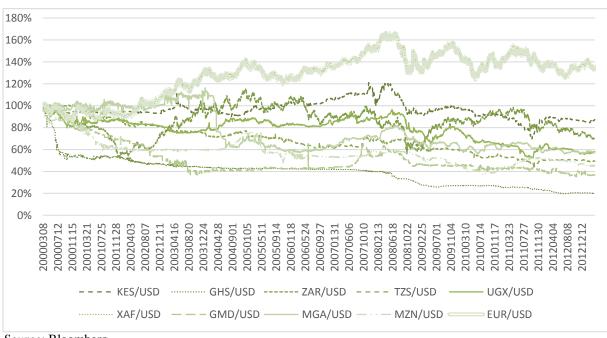
³ Czech Republic, Hungary, Poland, Slovakia

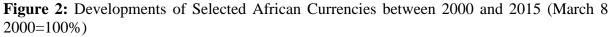
⁴ South African Rand, Brazil Real and Turkish Lira tended to be part of the same cluster in the past (Keskina et al. 2011).

 $^{^{5}} http://www.pbc.gov.cn/publish/english/955/2014/20140317160839706274217/20140317160839706274217_.html$

⁶Benin, Burkina-Faso, Côte d'Ivoire, Mali, Niger, Senegal, Togo, and Guinea-Bissau

⁷ http://go.worldbank.org/FKHEP1VQF0





Source: Bloomberg

CFA Franc (XAF) followed strictly the euro, due to its fixed regime, showing an appreciation against US dollar during the entire time set on Fig. 2. Kenyan Shilling (KES) and South African Rand (ZAR) presented an appreciating trend before the crisis only, otherwise all off the entire currency set depreciated against the US dollar – price-competitiveness improved from this aspect.

					normal distribution	autocorrelation	heteroscedasticity	stationarity
currency	mean	std	skewness	kurtosis	Jarque-Bera (p)	Ljung-Box (p)	ARCH-LM (p)	ADF (p)
KES/USD	0,00	0,01	0,32	20,62	0,00*	0,00**	0,22	0,00
GHS/USD	0,00	0,01	1,25	34,58	0,00*	0,00**	0,00***	0,00
ZAR/USD	0,00	0,01	1,05	15,75	0,00*	0,01**	0,14	0,00
TZS/USD	0,00	0,01	-0,83	40,58	0,00*	0,00**	0,00***	0,00
UGX/USD	0,00	0,01	0,47	17,08	0,00*	0,00**	0,02***	0,00
XAF/USD	0,00	0,01	-0,13	5,26	0,00*	0,00**	0,00***	0,00
GMD/USD	0,00	0,02	-0,05	174,94	0,00*	0,00**	0,03***	0,00
MGA/USD	0,00	0,01	1,73	59,42	0,00*	0,00**	0,00***	0,00
MZN/USD	0,00	0,01	0,81	43,05	0,00*	0,00**	0,00***	0,00
EUR/USD	0,00	0,01	-0,02	4,39	0,00*	0,51	0,59	0,00

Table 1: Descriptive Statistics of Currency Logarithmic Differentials

Notes: *: lack of normal distribution, **: autocorrelation at 2 lags, ***: heteroscedasticity at 2 lags, ****: unit root

Source: authors' calculations

Logarithmic differentials (1) of FX rates were tested to understand their basic characteristics on Tab. 1.

$$r_t = \ln(\frac{p_t}{p_{t-1}}) \tag{1}$$

Where r_t represents logarithmic return on trade day t, and p is FX rate.

Currencies had enormous fifth moments (kurtosis) compared to the ideal level of 3, suggesting that large magnitude fluctuations are more probable that they should be under the assumption of normal distribution. After this result the rejection of normal distribution by Jarque-Bera test was not a surprise. Time series were auto correlated with 2 lag, falsifying the common assumption about weak market efficiency. The appearance of heteroscedasticity suggests the future application of Generalized Autoregression Heteroscedasticity (GARCH) models to avoid correlation bias following Forbes and Rigobon (2002). There was no unit root in the data by the ADF test.

Contagions were defined as significant increase in correlations due to some kind of shock (2)

(2)

$\rho_{shock} \gg \rho_{tranqil}$,

where ρ_{shock} represents correlation under shock periods and $\rho_{tranqil}$ are correlations under "normal" periods. Current paper analyses two forms of shocks on the selected currency set i.e. a short-run and a long-run approach will be applied.

Short-run shocks were defined by non-normal distributed unconditional quantile of empirical data, referred as *fat tailed extreme returns* (r_{Xfat}) computed by the difference on the tails between theoretical normal and empirical distribution utilizing its "S"-shaped form, described by Clauset (2007) and Gabaix et al. (2003) to see the difference between theoretical and empirical returns under p_L low probability (3).

$$r_{\mathbb{X}fat+,p_L} \gg r_{normal,p_L} \text{ or } r_{\mathbb{X}fat-,p_L} \ll r_{normal,p_L} \text{ where } p_L \ll p_{E(r)}$$
(3)

Fat tailed extreme returns can appear both on negative (r_{Xfat-,p_L}) and positive (r_{Xfat+,p_L}) side of probability distribution under p_L low probable cases which are far from the probability of the expected value $(p_{E(r)})$.

Long-run shocks were defined by business cycles in the US and Eurozone following NBER⁸ and CEPR Euro Area Business Cycle Dating Committee⁹ data. Recession was defined by both of them as "a significant decline in economic activity spread across the economy, lasting more than a few months, normally visible in real GDP, real income, employment, industrial production, and wholesale-retail sales". US recession periods were defined by this method between March and November 2001 as well as between December 2007 and June 2009. Eurozone was in recession between January 2008 and April 2009 (which was really close to the US recession period) and after July 2011 and has not ended yet.

Different GARCH models were fitted on the data to manage heteroscedasticity of underlying data before later correlation fitting, following Cappeiello, Engle and Sheppard (2006). The applied GARCH(p,q), GJR GARCH(p,o,q), TARCH(p,o,q) and APARCH(p,o,q) (4-8) models can be useful to capture volatility developments and their clustering in time (heteroscedasticity).

GARCH (p,q):
$$\sigma_t^2 = \omega + \sum_{i=1}^p \alpha_i \varepsilon_{t-i}^2 + \sum_{j=1}^q \beta_i \sigma_{t-j}^2.$$
(4)

where σ_t^2 represents present variance, ω is a constant term, p denotes the lag number of squared past ε_{t-i}^2 innovations with α_i parameters, while q denotes the lag number of past σ_{t-j}^2 .variances with β_i parameters to represent volatility persistence. Asymmetric GARCH models can be introduced via

$$\begin{cases} S_{t-i}^{-} = 1, & \text{if } \varepsilon_{t-i} < 0 \\ S_{t-i}^{-} = 0, & \text{if } \varepsilon_{t-i} \ge 0 \end{cases} \text{ as a sign asymmetric reaction to decreasing returns.}$$
(5)

⁸ http://www.nber.org/cycles.html

⁹ http://www.cepr.org/content/euro-area-business-cycle-dating-committee

GJR GARCH (p,o,q):
$$\sigma_t = \omega + \sum_{i=1}^p \alpha_i |\varepsilon_{t-i}| + \sum_{i=1}^o \gamma_i S_{t-i}^- |\varepsilon_{t-i}| + \sum_{i=1}^q \beta_i \sigma_{t-i}, \qquad (6)$$

TARCH (p,o,q): $\sigma_t^2 = \omega + \sum_{i=1}^p \alpha_i \varepsilon_{t-i}^2 + \sum_{i=1}^o \gamma_i S_{t-i}^- \varepsilon_{t-i}^2 + \sum_{j=1}^q \beta_i \sigma_{t-j}^2 ,$ (7) $\label{eq:APARCH} \text{APARCH} (\text{p,o,q}) \text{:} \qquad \sigma_t^\delta = \omega + \sum_{i=1}^p \alpha_i (|\varepsilon_{t-i}| - \gamma_i \varepsilon_{t-i})^\delta + \sum_{j=1}^q \beta_j \sigma_{t-j}^\delta \;,$ (8)where $\alpha_i > 0$ (*i*=1,...,*p*), $\gamma_i + \alpha_i > 0$ (*i*=1,...,*o*), $\beta_i \ge 0$ (*i*=1,...,*q*), $\alpha_i + 0, 5 \gamma_j + \beta_k + <1$ (*i*=1,...,*p*, j=1,...,o, k=1,...,q) and δ index parameter can be between 1 and 2.

Model selection was made with focus on homoscedastic residuals (using a 2 lagged ARCH-LM test), searching for the lowest Bayesian Information Criteria (BIC). This study applies DCC-GARCH¹⁰ model, following Engle (2002), to analyze the daily common movements of the selected markets.

Short-run volatility-spillovers on GARCH residuals were analyzed by Vector Autoregression (VAR) models, following Rejeb and Boughrara (2015). Our VAR captures the dynamic interactions for a set of K vector of volatility series variables $h_t = (h_{1t}, ..., h_{Kt})'$. Substituting into the basic model of order p VAR has the form of (15) (Lütkepohl and Kratzig, 2004). $h_t = A_1 h_{t-1} + \dots + A_p h_{t-p} + u_t$ (9) Where the A_i 's are (*KxK*) coefficient matrices and $u_t = (u_{1t}, ..., u_{Kt})'$ is an unobservable error term - assumed to be a zero-mean independent white noise process with time-invariant, positive definite covariance matrix: $u_t \sim (0, E(u_t, u'_t))$. Our lag-selection focused on the lack of autocorrelation of u_t . Dynamic interactions among currency conditional volatilities (risk) depending on their own past values and other currencies tell many things about the interrelatedness of the selected currency.

The sample countries direct their foreign trade into the direction of US dollar (or USDpegged) and euro-denominated markets. A fixed exchange rate against euro endangers pricecompetitiveness under a period of euro-appreciation, while a floating regime is able to adapt. However, economic actors have to manage their FX exposure on both sides of their balance sheets, where a sudden change in currency common movement can undermine projectprofitability. For deeper understanding of the impact of shocks on common movements, contagion, divergence and interdependence were defined to capture all possible outcomes when we can find high levels of conditional correlations.

Contagion (10) occurs between *euro*, c_k currencies when the ρ^{euro,c_k} cross-market correlation becomes significantly higher due to a shock derived from EUR/USD market $(r_r^{EUR/USD})$ spreading to the others or as a result of other external factors (Forbes and Rigobon, 2002, Campbell et al., 2002, Bekaert et al., 2005):

$$r_x^{EUR/USD} = 1 \to \rho_n^{euro, c_k} \ll \rho_x^{euro, c_k},\tag{10}$$

Interdependence (11) occurs between *euro*, c_k currencies when the ρ^{euro,c_k} cross -market correlation is not significantly different, but the level of correlation is consistently high (Forbes and Rigobon, 2002): (11)

$$r_x^{EUR/USD} = 1 \rightarrow \rho_n^{euro,c_k} \approx \rho_x^{euro,c_k},$$

Divergence (12) occurs between *euro*, c_k currencies when the ρ^{euro,c_k} cross-market correlation becomes significantly lower due to a shock derived from one market $(r_r^{EUR/USD})$ spreading to others or as a result of other external factors (Bearce, 2002): $r_x^{EUR/USD} = 1 \rightarrow \rho_n^{euro,c_k} \gg \rho_x^{euro,c_k},$ (12)

¹⁰The estimation based on the Oxford MFE and UCSD toolboxs, developed by Kevin Sheppard: http://www.kevinsheppard.com/

This chapter summarized the available information on the data and the applied methodology to test, how selected African currencies behaved under the market shocks.

3 Results

The lack of normal distribution at the logarithmic returns suggested a success on the detection of fat tailed extreme fluctuations. Tab. 2 contains the results of the method, suggesting that the remaining r(n) truncated distribution converged closer to the ideal values of the first four moments: N~(0,1,0,3) – while the mean remained around zero, standard deviation was decreased (due to lack of excessive fluctuation), times series became less asymmetric (except GMD and EUR), while kurtosis is now closer to 3. At the same time, the applied method was able to manage the asymmetric appearance of extreme returns, while their overall mass remained lower than 10% in the entire data set.

currency		KES	GHS	ZAR	TZS	UGX	XAF	GMD	MGA	MZN	EUR
	entire	0,00	0,00	0,00	0,00	0,00	0,00	0,00	0,00	0,00	0,00
mean	r(n)	0,00	0,00	0,00	0,00	0,00	0,00	0,00	0,00	0,00	0,00
	entire	0,01	0,01	0,01	0,01	0,01	0,01	0,02	0,01	0,01	0,01
standard deviation	r(n)	0,00	0,01	0,01	0,00	0,00	0,00	0,01	0,01	0,01	0,00
	entire	0,32	1,25	1,05	-0,83	0,47	-0,13	-0,05	1,73	0,81	-0,02
skewness	r(n)	-0,08	0,09	-0,08	0,04	-0,02	-0,07	-0,07	0,03	-0,08	-0,05
	entire	20,62	34,58	15,75	40,58	17,08	5,26	174,94	59,42	43,05	4,39
kurtosis	r(n)	3,85	4,81	2,72	3,76	3,23	2,62	4,81	4,53	5,03	2,50
X+ treshold		0,01	0,02	0,02	0,01	0,01	0,01	0,04	0,03	0,03	0,01
X- treshold		-0,01	-0,02	-0,02	-0,01	-0,01	-0,01	-0,04	-0,03	-0,03	-0,01
Number of X-	+	118	101	162	84	114	192	73	68	83	220
Number of X-		102	102	73	75	82	149	53	64	66	163
Number of r(r	1)	3693	3710	3678	3754	3717	3572	3787	3781	3764	3530

Table 2: Fat Tailed Extreme Returns (Currencies in USD Denomination)

Notes: X+ means extreme positive fat tailed extreme returns, while X- means extreme negative fat tailed extreme returns. Usual or "normal" market days are represented by the r(n) truncated distribution. Highlighted area: fixed exchange rate policy.

Source: authors' calculations

The temporal distribution (percentage of extreme re) of fat tailed extreme returns was tested in Tab. 3, to check the increase of their mass under recession. US recession benchmark seemed to be better to capture a slightly increase in the occurrence of extreme fluctuations. The accommodative monetary policy in the Eurozone following the July 2011 crisis had a relaxing impact on currency pricing (except Ghana, Tanzania and Madagascar).

Table 3: Fat Tailed Extreme Returns (Currencies in USD Denomination)

	period		GHS	ZAR	TZS	UGX	XAF	GMD	MGA	MZN	EUR	recession
	March 2000 - February 2001	1%	20%	7%	2%	2%	6%	2%	4%	0%	6%	
	March 2001 - October 2001	9%	13%	10%	1%	1%	13%	26%	15%	3%	12%	"+"
	November 2001 - November 2007	9%	1%	9%	6%	5%	10%	2%	2%	4%	10%	
	December 2007 - May 2009	4%	0%	3%	3%	3%	9%	4%	1%	4%	12%	"+"
US	June 2009 - February 2015	4%	5%	4%	3%	3%	5%	3%	2%	1%	7%	
	March 2000 - December 2007	8%	4%	8%	5%	4%	10%	4%	4%	3%	10%	
	January 2008 - March 2009	4%	0%	2%	4%	3%	11%	5%	1%	4%	14%	''+''
	April 2009 - June 2011	7%	0%	6%	2%	5%	6%	3%	0%	0%	9%	
EU	July 2011 - March 2015	0%	11%	1%	4%	1%	2%	3%	4%	1%	4%	''+''

Source: authors' calculations

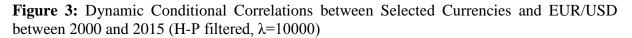
Heteroscedasticity of time series were managed by different GARCH models, where the selection was based on homoscedastic standardized residuals and lowest Bayesian Information Criteria (BIC). Alpha coefficient represented the importance of past innovations (with 1 or 2 lags), while gamma informs about asymmetric behaviour (higher volatility under depreciation), as well as beta refers to the volatility persistence (Tab. 4). Traditionally the volatility persistence has the highest value, so the result at GMD/USD is quite exceptional.

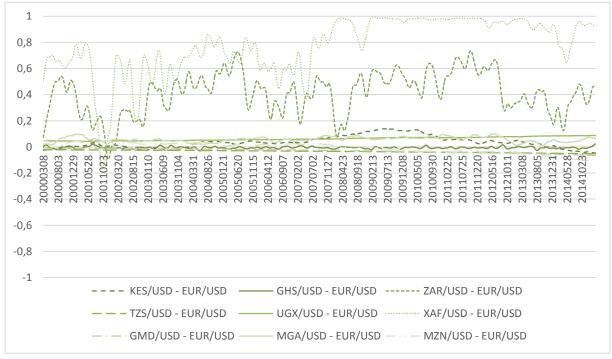
currency	model	constant	α_1	α2	γ	β_1	β_2	δ	BIC	
KES	TARCH(1,1,2)	0,00	0,23		0,04	0,44	0,32		-4,20	
GHS	GARCH(1,2)	0,00	0,11			0,37	0,51		-3,69	
ZAR	GJR GARCH(1,1,1)	0,00	0,02		0,10	0,93			-3,21	
TZS	GARCH(1,1)	0,00	0,20			0,80			-3,92	
UGX	GARCH(1,1)	0,00	0,19			0,81			-3,83	
XAF	GARCH(1,2)	0,00	0,06			0,20	0,91		-3,62	
GMD	TARCH(2,1,1)	0,00	0,14	0,28	0,17	0,49			-2,83	
MGA	APARCH(1,1,1)	0,00	0,02		0,06	0,95		3,44	-3,23	
MZN	GJR GARCH(1,1,2)	0,00	0,14		0,12	0,29	0,51		-3,32	
EUR	GARCH(1,1)	0,00	0,04			0,96			-3,71	

Table 4: GARCH Model Parameters (Currencies in USD Denomination)

Source: authors' calculations, UCSD toolbox

The Hodrick-Prescott (H-P) filtered (λ =10000) dynamic conditional correlation between euro and regional currencies proved to be strong only for XAF due to the pegged regime (Fig. 3). Other African floating currencies were uncorrelated, only the South African Rand (ZAR) showed some weak common movement. The H-P filter was applied to smooth the dataset, following Stavárek (2009).





Source: authors' calculations, UCSD toolbox

These results suggest weaker dependence, compared to the historically strong relations of Central-Eastern European currencies to euro Stavarek (2010) or Babetskaia-Kukharchuk et al. (2008). Therefore we can say that long-term correlation was not affected by business cycles in key economies (or by monetary responses on these developments). These results on Table 5 are suggesting that contagious processes can be checked only in South African and XAF currencies.

	period		GHS	ZAR	TZS	UGX	XAF	GMD	MGA	MZN	recession
	March 2000 - February 2001	-0,01	-0,01	0,21	- 0,03	0,05	0,54	-0,01	0,05	0,03	
	March 2001 - October 2001	-0,02	0,00	0,39	- 0,03	0,05	0,57	-0,01	0,05	0,02	"+"
	November 2001 - November 2007	0,05	-0,01	0,44	- 0,03	0,06	0,76	-0,02	0,06	0,01	
	December 2007 - May 2009	0,10	-0,01	0,51	- 0,04	0,07	0,98	-0,04	0,08	-0,03	" ₊ "
US	June 2009 - February 2015	0,02	-0,01	0,44	- 0,05	0,08	0,93	-0,05	0,05	-0,06	
	March 2000 - December 2007	0,04	-0,01	0,41	- 0,03	0,06	0,72	-0,02	0,05	0,02	
	January 2008 - March 2009	0,10	-0,01	0,50	- 0,04	0,07	0,98	-0,04	0,07	-0,03	"+"
	April 2009 - June 2011	0,04	-0,01	0,53	- 0,04	0,07	0,97	-0,05	0,07	-0,04	
EU	July 2011 - March 2015	-0,01	-0,01	0,34	- 0,05	0,08	0,88	-0,06	0,04	-0,08	"+"

Table 5: Average Dynamic Conditional Correlation (Currencies in USD Denomination)

Source: authors' calculations, UCSD toolbox

Short-term currency market developments were captured in several cases when extreme and normal subsets of dynamic conditional correlations (DCCs) were compared with two-sided t-tests to capture contagions, divergences or interdependence (Tab. 6). South African Rand (ZAR) and the EUR-pegged XAF had lost their strong common movement under turbulent times, presenting a divergence – and a vulnerability of the monetary framework.

Table 6: Difference between Highest Dynamic Conditional Correlations

	ZAR/USD-EUR/USD	XAF/USD-EUR/USD
avg. correlation(n)-avg. correlation(x+)	-0,04**	-0,02**
avg. correlation(n)-avg. correlation(x-)	-0,03**	0,00
		,,

Notes: *: contagion, **: divergence

Source: authors' calculations, UCSD toolbox

The short-term volatility spillover was been studied with VAR models (Fig. 4) on the vectors of volatility series to see, how daily conditional standard deviations are related to each other. Some currencies seemed to be isolated from other's volatility, like KES, GHS, MGA while others like ZAR, UGX, GMD interfered with another sub-Saharan currencies. Only TZS, XAF, MZN presented significant beta coefficients for EUR – for XAF this behaviour was supported by its EUR-peg.

Figure 4: Vector Autoregressions on Conditional Standard Deviations (currencies in USD denomination)

 $h_{t,KES} = 0.6960 h_{t-1,KES} + 0.3292 h_{t-2,KES} - 0.0671 h_{t-4,KES}$, R²=0.88, lag no. = 4 $h_{t,GHS} = 0.7138h_{t-1,GHS} + 0.3625h_{t-2,GHS} - 0.0942h_{t-3,GHS}$, R²=0.96, lag no. = 3 $h_{t,ZAR} = 0.0173h_{t-1,KES} + 0.9749h_{t-1,ZAR} + 0.0740h_{t-1,TZS}$, R²=0.97, lag no. = 1 $h_{tTZS} =$ $0.0511 h_{t-1,KES} + 1.1349 h_{t-1,TZS} - 0.2024 h_{t-2,TZS} + 0.2809 h_{t-1,EUR} - 0.2340 h_{t-2,EUR} + 0.000 h_{t-1,EUR} +$ 0.0002, R²=091, lag no. = 2 $h_{t,UGX} = 0.0300h_{t-2,GHS} + 1.046h_{t-1,UGX} - 0.1208h_{t-2,UGX} + 0.0002$, R²=0.89, lag no. = 2 $h_{t,XAF} =$ $-0.0129 h_{t-2,KES} + 0.0097 h_{t-1,TZS} + 0.1571 h_{t-1,XAF} + 0.9050 h_{t-2,XAF} - 0.0965 h_{t-3,XAF} + 0.0097 h_{t-1,TZS} + 0.0097$ $0.0203h_{t-1,MGA} - 0.0215h_{t-2,MGA} + 0.0089h_{t-2,MZN} - 0.0113h_{t-3,MZN} - 0.1847h_{t-1,EUR} + 0.0089h_{t-2,MZN} - 0.0113h_{t-3,MZN} - 0.01847h_{t-1,EUR} + 0.0089h_{t-2,MZN} - 0.00113h_{t-3,MZN} - 0.011847h_{t-1,EUR} + 0.0089h_{t-2,MZN} - 0.00113h_{t-3,MZN} - 0.0011847h_{t-1,EUR} + 0.0089h_{t-2,MZN} - 0.00118h_{t-3,MZN} - 0.0011847h_{t-1,EUR} + 0.0089h_{t-2,MZN} - 0.00118h_{t-3,MZN} - 0.0018h_{t-3,MZN} - 0.0018h_{t-3,MZN} - 0.0018h_{t-3,MZN} - 0.0018h_{t-3$ $0.2000h_{t-2.EUR} + 0.000$, R²=0.96, lag no. = 3 $h_{t.GMD} =$ $0.3740h_{t-3,GMD} - 0.1463h_{t-4,GMD} + 0.0634h_{t-4,GMD} + 0.0022$, R²=0.76, lag no. = 5 $h_{t,MGA} = 1.064 h_{t-1,MGA} - 0.0606 h_{t-3,MGA}$, R²=0.98, lag no. = 3 $h_{t,MZN} =$ $0.7137h_{t-1,MZN} + 0.3675h_{t-2,MZN} - 0.0518h_{t-3,MZN} + 0.0477h_{t-4,MZN} - 0.1143h_{t-5,MZN} + 0.0477h_{t-4,MZN} - 0.01143h_{t-5,MZN} + 0.0477h_{t-4,MZN} + 0.0477h_{t$ $0.7086h_{t-1,EUR} - 1.1576h_{t-2,EUR} + 0.0005$, R²=0.93, lag no. = 5 $h_{t,EUR} = 0.003h_{t-1,KES} + 0.002h_{t-1,TZS} + 0.003h_{t-1,UGX} + 0.012h_{t-1,XAF} + 0.9812h_{t-1,EUR},$ $R^2 = 0.99$, lag no. = 1

Source: authors' calculations, JPL toolbox

The African floating currencies presented some special features in their short-term dynamics: most of them were uncorrelated with EUR/USD rate despite the strong trade and political ties but monetary easing after 2008 decreased the occurrence of extreme currency fluctuations. The short-term volatility spillover supported the idea of market isolation, because currency volatility was affected only by its own past values.

Conclusion

Current article evaluated the FX exposure of the African countries with floating exchange rate regimes to compare them with the pegged alternative, like XAF. The analysed countries have many bounds to the developed countries: among ordinary trade relations, their integration into the world economy is promoted via preferential tariff agreements, aid and discounted credit programs. The economic actors in the sample countries had to pay the price of floating regimes as the results about the fat-tailed extreme fluctuations and volatility models suggested, but their currencies were uncorrelated with euro and mostly isolated from its volatility as well. Therefore contagions were not able to emerge on these markets despite the real, political and financial links among these countries and the Eurozone or US market, but the flexibility of the FX regimes allowed the depreciation against euro to maintain price-competitiveness. It means that international and national development and aid programs shall not focus on the problem at the sample currencies are not attractive enough to step on the menu of international investors. After a currency enters to the "emerging" status like CEE or

South African foreign exchanges, their movements will be influenced by international portfolio re-balancing decisions.

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